Statistical Arbitrage Opportunities Using Bellman Equation

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Abstract. This paper uses the dynamic programming to detect the optimal statistical arbitrage opportunities in a market including a bond and a stock. First, it is assumed that the growth rates of stock are independent random variables and Bellman equation is derived for probability of gain of a portfolio containing a long position in stock and short position in bond. The Bellman equation is derived and its approximations are studied. Then, using the simulation, the performance of method in correlated growth rates cases is proposed. Conclusions are also given.

Keywords: Bellman equation; Bond; Dynamic programming; Probability of statistical arbitrage; Stock

1 Introduction. This paper deals with construction of statistical arbitrage strategy to obtain possible profit with a high probability in future with zero investment at the current time. Jarrow *et al*. (2005) investigated the relationship of market efficiency and statistical arbitrage. A comprehensive review in statistical arbitrage may be found in Pole (2007). Avellaneda and Lee (2008) applied the statistical arbitrage detection techniques in U.S. equities markets. In foreign exchange markets, there are many basic shortest path algorithms to detect statistical arbitrage, see Goldberg (2008). There are many basic strategies to this end, like market neutral strategies, pairs trading, high frequency trading, co-integration trading and mean revert trading, see Aldridge (2009) and references therein. The dynamic programming is a method for optimal control like (i) maximum principle and (ii) calculus of variation which may be used to exploit the statistical arbitrages, see Intriligator (2002) and Zisserman (2013).

Here, it is assumed that B_t is a bond with riskless rate r_f growth, that is $B_t = B_0(1+r_f)^t$ and S_t is a stock with growth rate $r_t, t \geq 1$, i.e.,

$$S_t = S_0 \prod_{i=1}^{t} (1+r_i).$$

In this paper, is it assumed that r_i's are independent random variables and for each $t \geq 1$, $E(r_t) > r_f$. The value of a portfolio with long position in one unit of stock and α_t unit of bond in short position is $\pi_t = S_t - \alpha_t B_t$. Suppose that $\pi_0 = 0$, then $\alpha_0 = \frac{S_0}{B_0}$. For $t \geq 1$, let $\gamma_t = \ln\left(\frac{\alpha_t}{\alpha_0}\right)$ and p_t be the gain probability that is $p_t = P(\pi_t > 0)$. Then,

$$p_t = P(S_t - \alpha_t B_t > 0) = P\left(\frac{S_t}{B_t} > \alpha_t\right) =$$

$$P\left(\sum_{i=1}^{t}\left[\ln(1+r_i) - \ln(1+r_f)\right] > \gamma_t\right) \approx P\left(\sum_{i=1}^{t}(r_i - r_f) > \gamma_t\right).$$

The last approximation is true because $\ln(1+x) \approx x$ for small x. Let F_t be the distribution function of $\sum_{i=1}^{t}(r_i - r_f)$. Then, $p_t = 1 - F_t(\gamma_t)$. Notice that, $p_t = p_{t-1} - F_t(\gamma_t) + F_{t-1}(\gamma_{t-1})$. One can see that

$$F_t(x) = E\{F_{t-1}(x + r_f - r_t)\}.$$

Assuming, $F_{t-1}(\gamma_t) - F_{t-1}(\gamma_{t-1}) \approx 0$, then,

$$p_t = p_{t-1} + E\{F_{t-1}(\gamma_t) + F_{t-1}(\gamma_t + r_f - r_t)\}.$$

Therefore, to maximize the gain probability with respect to γ_t, it is enough to solve the following

equation

$$p_t = \max_{\gamma_t} (p_{t-1} + E\{F_{t-1}(\gamma_t) + F_{t-1}(\gamma_t + r_f - r_t)\}).$$

The last equation defines a dynamic programming framework for p_t. As soon as γ_t is determined, then $\alpha_t = \alpha_0 e^{\gamma_t}$ is derived. The rest of paper is organized as follows. In section 2, assuming normal distribution for r_t's the p_t, γ_t and α_t are derived. The exact and approximated solutions are derived. Section 3 examines the correlated r_t's. A real data set is analyzed.

2 Normal r_t's. In this section, first, the exact dynamic programming is presented. Then, the approximated solution is proposed. To this end, suppose that r_t's have normal distribution with mean μ_t and variance σ_t^2.

2.1 Exact solution. The F_t is distribution function of normal distribution with mean $\sum_{i=1}^{t}(\mu_i - r_f)$ and variance $\sum_{i=1}^{t}\sigma_i^2$. For two independent normally distributed random variables X and Y, it is true that $E(F_X(Y)) = P(X \le Y)$, thus,

$$E\{F_{t-1}(\gamma_t + r_f - r_t)\} = F_t(\gamma_t),$$

then,

$$p_t = \max_{\gamma_t} (p_{t-1} + F_{t-1}(\gamma_t) + F_t(\gamma_t)).$$

Example 1. Let $r_f = 0.05$ and $\mu_i = a\mu_{i-1}$ for $|a| < 1$ and $\sigma_t^2 = 0.16$. Notice that $p_0 = 0$. Let $\alpha_0 = 1$. The following Table gives the values of p_t, γ_t and α_t for $a = 0.2$.

Table 1: The values of p_t, γ_t and α_t

t	1	2	3	4	5	6	7	8	9	10
p(t)	0.172	0.22	0.234	0.239	0.351	0.36	0.367	0.373	0.378	0.383
α(t)	0.27	0.48	0.144	8.891	0.105	0.079	0.089	0.081	0.074	0.081
γ(t)	1.305	0.73	1.94	-2.185	2.26	2.54	2.425	2.515	2.61	2.5

2.2 Approximations. The following three propositions considers different approximations. To study the approximated solution, notice that

$$p_t - p_{t-1} = F_{t-1}(\gamma_{t-1}) - F_t(\gamma_t).$$

Proposition 1. For special case for $\gamma_t = -bt$, for some $b > 0$.

(a) The optimum b is

$$b = \frac{tF'_{t-1}(0) - (t-1)F'_t(0)}{t^2F''_{t-1}(0) + (t-1)^2F''_t(0)}.$$

(b) For fixed t, the optimum b is given by Newton-Raphson recursive formulae given by

$$b_{k+1} = b_k - \frac{g(b_k)}{g'(b_k)},$$

where $g(x) = tF'_t(-tx) - (t-1)F'_{t-1}(-(t-1)x)$.

Proof. Notice that

$$p_t - p_{t-1} = \max_{\mathrm{b}} \{F_{t-1}(0) - btF'_{t-1}(0) + \frac{b^2t^2F''_{t-1}(0)}{2} - F_t(0) + b(t-1)F'_t(0) + \frac{b^2(t-1)^2F''_t(0)}{2}\}.$$

This completes the proof of part (a). For part (b), notice that the optimum b, is obtained by letting the derivative of $F_{t-1}(\gamma_{t-1}) - F_t(\gamma_t)$ with respect to b equal to zero.

Proposition 2. If $\gamma_t - \gamma_{t-1} = hd_t$, then

$$p_t - p_{t-1} = F_{t-1}(\gamma_t) - F_t(\gamma_t) - hd_t F'_{t-1}(\gamma_t) + \frac{h^2 d_t^2}{2} F''_{t-1}(\gamma_t).$$

Proof. Use the straightforward application of Taylor series completes the proof. The following proposition studies the Monte Carlo approximation.

Proposition 3. The Monte Carlo approximation for $p_t - p_{t-1}$ is given by

$$\frac{1}{M}\sum_{i=1}^{M}(F_{t-1}(\gamma_t) - F_{t-1}(\gamma_t + r_f - r_t^i),$$

where $r_t^i, i = 1,2,\dots,M$ is a Monte Carlo sample from r_t.

3 Correlated r_t's. In financial applications, usually, it is not true to assume that r_t's are independent. In this section, it is assumed that they follow a first order autoregressive process with GARCH(1,1) errors, a rich practical model which works for many types of financial data sets, that is

$$r_t - r_f = \alpha + \beta(r_{t-1} - r_f) + \sigma_t z_t,$$

where z_t's are iid standard normal random variables and σ_t constitutes a GARCH(1,1) series. In the rest of this section, the existence the statistical arbitrage in stock of Intel corporation, a multinational technology company, is surveyed. The daily stock price are collected for period of study 20 Feb 2015 to 18 Feb 2016, including 250 log-returns. They are taken from Google-finance website[1]. Time series plot of return series is given by Figure 3.1. The x-axis is time and y-axis is the logarithmic return.

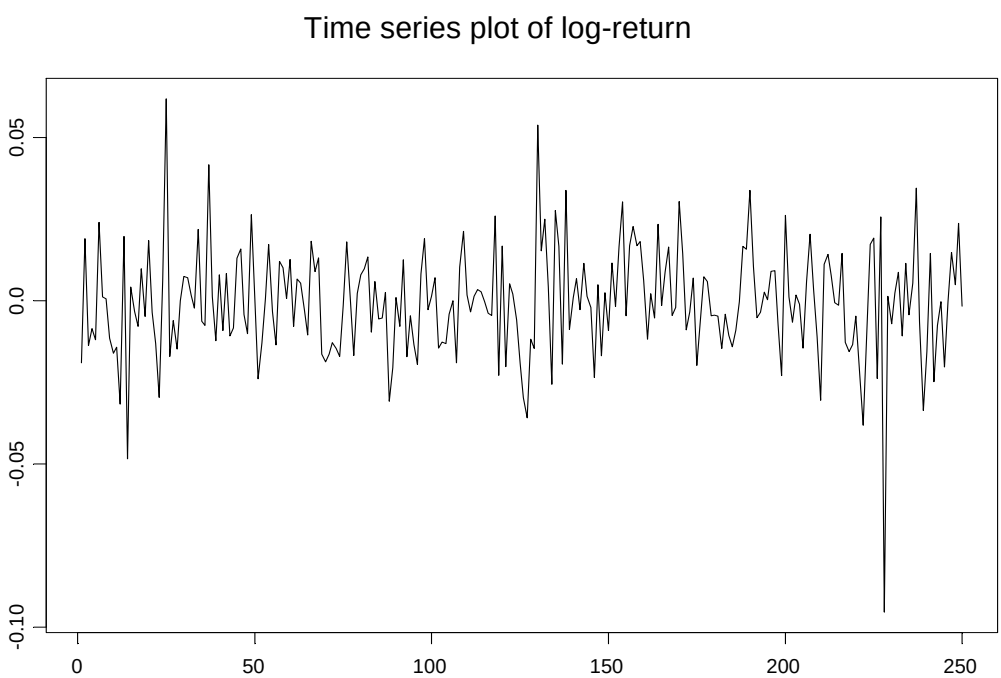

Figure 3.1: Time series plot of log-returns

[1] http://www.google.com/finance?cid=284784

According to the Proposition 1 part (a), for $\gamma_t = -bt$, the optimum b is

$$b = \frac{tF'_{t-1}(0) - (t-1)F'_t(0)}{t^2F''_{t-1}(0) + (t-1)^2F''_t(0)}.$$

The quantities of $F'_{t-1}(0)$, $F'_t(0)$, $F''_{t-1}(0)$ and $F''_t(0)$ are derived using the Monte Carlo approximation and numerical differentiating. The plot of p_t is given as follows in Figure 3.2. The x-axis is time and y-axis is the probability of statistical arbitrage.

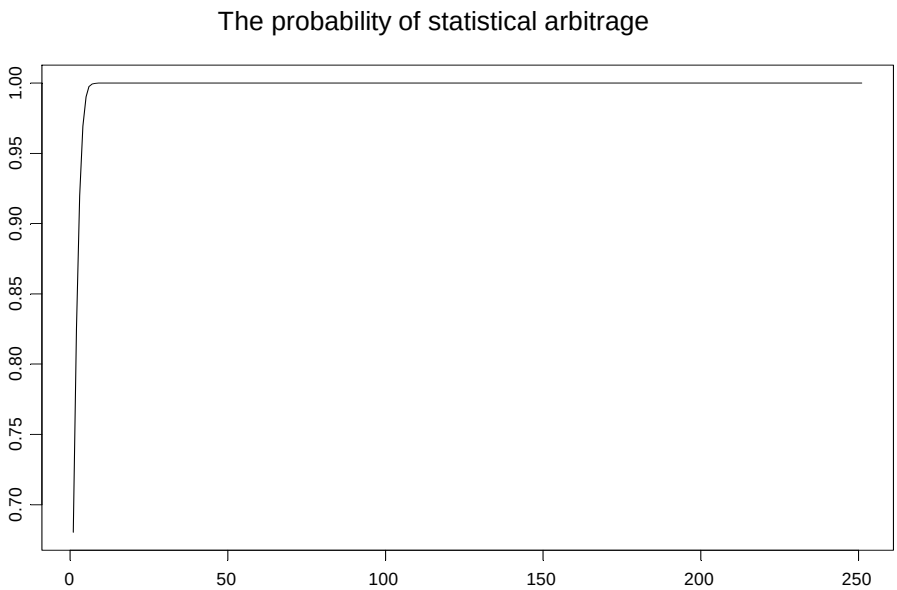

Figure 3.2: The probability of statistical arbitrage over time

4. Conclusions. The application of Bellman equation is studied to fine the maximum probability of achieving the statistical arbitrage opportunities. Two cases of independent and correlated returns are studied and in both cases high probability of statistical arbitrage is derived.

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