

Mathematical Programming for Statistical Inference**Abstract**

Mathematical programming as a non-classical method arises as a tool for finding the global optimal solution. The aim of this paper is to suggest mathematical programming models for hypothesis tests, the researcher make balance between the high power ($1-\beta$), and the probability of a type I error, significance (α), of the test. The paper introduces a simulation study to evaluate the performance of the two suggested mathematical programming model for tests hypotheses. The two suggested mathematical programming models solved with different sample sizes and different level of significant. The suggested model calculates the critical values which determine the rejection region exactly and the results are easy to interpret clearly. Then the conclusion for the suggested mathematical programming models makes balance between the power and the significance.

Key words:

Hypothesis Tests, Power of a statistical test, Significant, Mathematical Programming, Type I and Type II errors.

1. Introduction

In almost 300 years since its introduction by Arbuthnot (1710), null hypothesis significance testing (NHST) has become an important tool for working scientists. Since then it has expanded into one of the most widely used quantitative methodologies, and has found its way into nearly all areas of human endeavour [6].

Almost the true value of population parameters was unknown. They must be estimated. The major purpose of hypothesis testing is to choose between two competing hypotheses about the value of a population parameter.

The power of a statistical test is $(1-\beta)$, which is the probability of rejecting H_0 when H_0 is false. **The level of significance** is the maximum probability of committing a type I error. Usually $\alpha = 0.01, 0.05$ or 0.10 which means there are 90%, 95% or 99% confident in rejecting the null. **The critical value** separates the critical region from the noncritical region. The symbol for critical value is C.V. **The critical or rejection region** is the range of values of the test value that indicates that there is a significant difference and that the null hypothesis should be rejected. **The noncritical or non-rejection region** is the range of values of the test value that indicates that the difference was probably due to chance and that the null hypothesis should not be rejected [7], [3] and [8].

When testing hypotheses, the researcher would like the test to have high power which means the ability to conclude the null hypothesis is false when it really is false with high probability. And also like the probability of a type I error, α , of the test to be small. Unfortunately, making the probability of a type I error smaller makes the test less powerful; making the test more powerful leads to a higher type I error. Therefore, a compromise is needed between these competing goals when performing hypothesis testing [1].

The value of β is not easy to compute. It depends on several things, including the value of α , the size of the sample, the population standard deviation, and the actual difference between the hypothesized value of the parameter being tested and the true parameter. The researcher has control over two of these factors, namely, the selection of α and the size of the sample. The major problem, then, lies in knowing the actual difference between the hypothesized parameter and the true parameter. If this difference were known, then the value of the parameter would be known; and if the parameter were known, then there would be no need to do any hypothesis testing. Hence, the value of β cannot be computed. But this does not mean that it should be ignored [3].

The power of a statistical test measures the sensitivity of the test to detect a real difference in parameters if one actually exists. The power of a test is a probability and, like all probabilities, can have values ranging from 0 to 1. In other words, the closer the power of a test is to 1, the better the test is for rejecting the null hypothesis if the null hypothesis is, in fact, false.

The power of a test can be increased by increasing the value of α . Recall that as α increases, β decreases. So if β is decreased, then $1-\beta$ will increase, thus increasing the power of the test. Another way to increase the power of a test is to select a larger sample size. A larger sample size would make the standard error of the mean smaller and consequently reduce β . The relationship among α , β , and the power of a test can be analyzed in greater detail than the explanation given here. The goal, then, is to try to keep the probabilities of type I and type II errors as small as possible [3].

Hypothesis testing is the most widely used method for statistical inference in the world. It forms the basis for decision science and all business-oriented decision-making methods. Yang *et al* [11] proposed a time-varying effect model that can be used to characterize gender-specific trajectories of health behaviours and conduct hypothesis testing for gender differences. In terms of the performance of the hypothesis testing, the type I error rates are close to their corresponding significance levels under all combinations of sample size and number of time points. Furthermore, the power increases as the alternative hypothesis deviates more from the null hypothesis, and the rate of this increasing trend is higher when the sample size and the number of time points are larger.

Luo and Yang [9] presented hypothesis testing is adopted to solve the water pollution detection problems. Firstly, a brief description of the monitoring sensor network and what problems there are in the pollution detection are given. Secondly, theoretical approaches to solve the detection problems are analyzed based on hypothesis testing. Thirdly, the specific detection algorithms are given. Finally, implementation examples are given to illustrate the proposed pollution detection methods.

Chen *et al* [4] developed a mathematical programming model to construct the $(1 - \alpha) \times 100\%$ confidence interval of the process capability index. They examined the effectiveness of the proposed approach by using Monte Carlo simulation to verify the reliability of the proposed model by assessing coverage percentage for 95% confidence interval. The results indicated that deriving the upper and lower limits of the confidence interval eliminated the need for the complex computation processes encountered in statistical methods and it made it possible for enterprises to evaluate process performance easily and effectively.

This paper included three additional sections: the mathematical programming models formulation of hypotheses tests is presented in section 2. In section 3 introduced a simulation study to evaluate the performance of the two suggested mathematical programming model for tests hypotheses. Finally, concluding remarks are provided in Section 4.

2. Mathematical programming models for tests hypotheses

- Neyman and Pearson 1936 determining a critical region such that for a given level of significance, the second kind of error is minimized or equivalently the power function is maximized is known as the Neyman-Pearson problem. Neyman and Pearson also established a lemma which is of fundamental importance in the theory of testing statistical hypotheses.

Arthanari and Dodge [2] stated theoretical mathematical programming for Neyman-Pearson problem as:

$$\begin{aligned} \text{Maximize} \quad & 1 - P[S - W | H_1] \quad \xrightarrow{\text{power}} (1-\beta) \quad (1) \\ \text{s.t} \end{aligned}$$

$$P[W|H_0] = \alpha \text{ (A given value)} \quad (2)$$

where H_0 is a null hypothesis, H_1 is an alternative hypothesis,

β is type II error, α is level of significance or type I error,

S is a sample space, where $(S - W)$ is a non-rejection region and W is a rejection region.

- When simple H_0 against simple H_1 with finite sample space Arthanari and Dodge [2] equivalent the previous theoretical mathematical programming for Neyman-Pearson problem as following:

$$\text{Maximize} \quad 1 - \sum P_{1i}(1 - \phi_i) \quad (3)$$

s.t

$$\sum P_{0i}\phi_i = \alpha \quad (4)$$

$$\phi_i = 0 \text{ or } 1 \quad \text{for all } i \in I^*$$

where $S = \{x_1, x_2, \dots, x_s\}$, $I^* = \{1, 2, \dots, s\}$ be the set of all indices of $x \in S$.

And ϕ_i be the indicator of any $W \subset S$, that is,

$$\phi_i = \begin{cases} 1 & \text{if } i \in W \\ 0 & \text{otherwise.} \end{cases}$$

where $\phi = (\phi_1, \dots, \phi_s)$, P_{0i} the probability that the realization is x_i given that H_0 is true. Similarly P_{1i} is defined, where α is level of significance or type I error and the summation is over all $i \in I^*$.

This problem can be readily being seen as a linear-programming problem with the additional restriction that ϕ_i is 0 or 1. Such problems are known as 0-1 integer programming problems especial when model have only one constraint.

The study formed and solved the previous model as mathematical programming model in the following section.

- The hypotheses tests applicable mathematical programming model depended on the previous models is suggested. The suggested hypotheses tests mathematical programming model applied when Binomial distribution used as probability distribution as follows:

$$\text{Maximize } P[x \geq K | P_1] = \sum_{i=k}^n \binom{n}{i} p_1^i (1-p_1)^{n-i} \quad (5)$$

$$\text{s.t.} \\ P[x \geq K | P_1] = \sum_{i=k}^n \binom{n}{i} p_0^i (1-p_0)^{n-i} \leq \alpha \quad (6)$$

$$0 \leq k \leq n$$

where $x \sim \text{Bin}(n, p)$.

Johnson *et al* [5] showed that there is a relation between cumulative Binomial distribution and incomplete Beta distribution. The relation is as follows:

$$\sum_{i=k}^n \binom{n}{i} p^i (1-p)^{n-i} = \frac{1}{B(k, n-k+1)} \int_0^p t^{k-1} (1-t)^{n-k} dt \quad (7)$$

The standard discrete distributions may arise as a result of the sampling scheme adopted. In sampling with replacement n items from a lot of manufactured items, the number of defectives follows the binomial law. On the other hand, if one uses what is known as the inverse binomial sampling procedure (that is, one goes on sampling with replacement until he get a fixed number such as k of defectives), the number of items sampled follows the negative binomial distribution.

Let

$$y(x, p, k) = \binom{k+x-1}{x} p^k (1-p)^x \quad (8)$$

where $0 < p < 1, 0 < k < \infty, x = 0, 1, 2, \dots$

In order to evaluate the negative binomial distribution function:

$$Y(r, p, k) = \sum_{x=0}^r y(x, p, k) \quad (9)$$

positive binomial distribution function tables can be used, when k is a positive integer, or tables of the incomplete beta function for any k . Thus no special tables are needed, since:

$$Y(r, p, k) = 1 - B(k-1, p, r+k), k = 1, 2, 3, \dots \quad (10)$$

where:

$$B(c, p, n) = \sum_{x=0}^c \binom{n}{x} p^x (1-p)^{n-x} \quad (11)$$

and

$$Y(r, p, k) = I_p(k, r+1) \quad 0 < k < \infty \quad (12)$$

where:

$$I_P(m, n) = \frac{1}{B(m, n)} \int_0^p u^{m-1} (1-u)^{n-1} du \quad (13)$$

(the complete beta function) , [10].

Then the suggested hypothesis test mathematical programming model became as follows:

$$\text{Maximize } I_t(k, n-k+1) = \frac{1}{B(k, n-k+1)} \int_0^{p_1} t^{k-1} (1-t)^{n-k} dt \quad (14)$$

s.t

$$\frac{1}{B(k, n-k+1)} \int_0^{p_0} t^{k-1} (1-t)^{n-k} dt \leq \alpha \quad (15)$$

$$0 \leq k \leq n$$

The study formed and solved this model as mathematical programming model in the following section.

3. The Simulation study

This section introduces a simulation study to evaluate the performance of the two suggested mathematical programming model for tests hypotheses. The two suggested mathematical programming model determinate power of a statistical test and the critical value K. The simulation study takes two different phases as following:

First phase:

The first phase in the simulation study used Arthanari and Dodge [2] theoretical mathematical programming for Neyman-Pearson for tests hypotheses. It takes steps as following:

1. The study used three different sample sizes 6, 10 and 35 as an initial step.
2. The study used $\alpha = 0.05$.
3. P_0 and P_1 are the probability of the binomial distribution.
4. P_0 and P_1 were chosen arbitrarily.
5. The study interested to calculate K which maxims the power with different P_0 and P_1 to all sample sizes (6, 10 and 35).
6. These simulated runs have been done through building routines using GAMS statistical packages.
7. GAMS statistical packages were used to solve the Arthanari and Dodge [2] theoretical mathematical programming model for Neyman-Pearson for tests hypotheses.

Table 1 the results of the first suggested model.

n	P ₀	P ₁	Number of reject K	Power of test
6	0.2	0.3	2	0.969
	0.2	0.1	1	0.9995
	0.3	0.7	2	0.472
	0.3	0.4	2	0.913

	0.4	0.3	1	0.998
	0.4	0.5	1	0.969
	0.5	0.6	1	0.922
	0.6	0.7	1	0.998
	0.7	0.3	2	0.472
	0.7	0.8	2	0.969
10	0.2	0.3	4	0.901
	0.2	0.1	2	1
	0.4	0.5	4	0.9908
	0.5	0.6	4	0.926
	0.6	0.7	3	0.955
	0.7	0.8	2	0.865
	0.9	0.8	4	0.804
35	0.2	0.3	17	0.651
	0.4	0.5	16	0.727
	0.6	0.7	21	0.737

The study used three different sample sizes (6, 10 and 35) with P_0 , P_1 and $\alpha = 0.05$. Table 1 showed the results when $n=6$ the power of the test approximately one and the number of reject between 1 or 2 with different P_0 and P_1 . The results when $n=10$ the power of the test approximately one and the number of reject between 2 to 4 with different P_0 and P_1 . The results when $n=35$ the power of the test approximately one and the number of reject between 16 to 21 with different P_0 and P_1 .

The model solutions appear as 0-1 integer programming problems. It determined critical or rejection region as points. It rejected some points and accepted some points. That can be difficult to determine the critical or rejection region exactly and the results are difficult to interpret clearly.

Second phase:

The second phase in the simulation study used the suggested hypotheses tests mathematical programming model which shown as previous.

The second phase of the simulation study takes steps as following:

1. The study interested to calculate K which maxims the power with different α , P_0 and P_1 to all sample sizes (10, 20, 30, 40 and 50).
2. The study used different vales of $\alpha = 0.01$, 0.05 and 0.10.
3. P_0 and P_1 are the probability of the binomial distribution which chosen arbitrarily.
4. P_0 and P_1 were chosen when different decimal (one decimal, two decimal and three decimal) are used.

5. The study used five different sample sizes, the first 10, 20 which represent small sample sizes, the second is 30,40 which moderate sample sizes while the third is 50 which represent more larger than 10, 20.
6. These simulated runs have been done through building routines using GAMS statistical packages.
7. GAMS statistical packages were used to solve the suggested mathematical programming model for hypotheses tests.

Table 2 the second suggested model with $\alpha= 0.01$

P₀	P₁	n = 10		n = 20		n = 30		n = 40		n = 50	
		k	Power of test	k	Power of test	k	Power of test	k	Power of test	k	Power of test
0.3	0.7	1	1	3	1	5	1	8	1	8	1
0.4	0.5	2	0.989	5	0.994	9	0.992	12	0.997	16	0.997
0.42	0.53	2	0.994	6	0.989	9	0.997	13	0.997	17	0.998
0.442	0.551	2	0.996	6	0.994	10	0.995	14	0.997	18	0.998
0.4	0.6	2	0.998	5	1	9	1	12	1	16	1
0.4	0.7	2	1	5	1	9	1	12	1	14	1
0.5	0.6	3	0.988	7	0.994	11	0.997	16	0.997	20	0.999
0.52	0.67	3	0.997	8	0.997	12	0.999	17	1	21	1
0.535	0.681	3	0.998	8	0.998	13	0.999	17	1	22	1
0.5	0.7	3	0.998	7	1	11	1	2	1	20	1
0.6	0.7	4	0.989	9	0.995	15	0.994	20	0.998	26	0.986

The study interested to calculate K which maxims the power with different α , P_0 and P_1 to all sample sizes (10, 20, 30, 40 and 50). Table 2 showed the results when different sample sizes, and $\alpha = 0.01$ the power of the test approximately one and the number of reject K between 1 to 26 with different sample sizes and different P_0 and P_1 .

Table 3 The second suggested model with $\alpha=0.05$

P_0	P_1	n = 10		n = 20		n = 30		n = 40		n = 50	
		k	Power of test	k	Power of test	k	Power of test	k	Power of test	k	Power of test
0.3	0.7	1	1	3	1	5	1	7	1	10	1
0.4	0.5	2	0.989	4	0.999	8	0.997	11	0.999	14	1
0.42	0.53	2	0.994	5	0.997	8	0.999	12	0.999	15	1
0.442	0.551	2	0.996	5	0.999	8	0.998	13	0.999	16	1
0.4	0.6	2	0.998	4	1	8	1	11	1	14	1
0.4	0.7	2	1	4	1	8	1	8	1	13	1
0.5	0.6	2	0.998	6	0.998	11	0.997	15	0.999	19	0.999
0.52	0.67	3	0.997	7	0.999	11	1	16	1	20	1
0.535	0.681	3	0.998	7	0.999	12	1	16	1	21	1
0.5	0.7	2	1	6	1	2	1	15	1	19	1
0.6	0.7	3	0.998	8	0.999	14	0.998	19	0.999	24	0.997

Table 3 showed the results when different sample sizes, when $\alpha=0.05$ the power of the test approximately one and the number of reject between 1 to 24 with different sample sizes and different P_0 and P_1 .

Table 4 The second suggested model with $\alpha=0.10$

P_0	P_1	n = 10		n = 20		n = 30		n = 40		n = 50	
		k	Power of test	k	Power of test	k	Power of test	k	Power of test	k	Power of test
0.3	0.7	1	1	2	1	3	1	6	1	8	1
0.4	0.5	1	0.999	3	1	6	1	9	1	12	1
0.42	0.53	1	0.999	3	1	7	1	10	1	13	1
0.442	0.551	1	1	4	1	7	1	11	1	14	1
0.4	0.6	1	1	3	1	2	1	9	1	12	1
0.4	0.7	1	1	1	1	2	1	8	1	10	1
0.5	0.6	1	1	5	1	9	1	13	1	17	1
0.52	0.67	2	1	5	1	9	1	13	1	18	1
0.535	0.681	2	1	6	1	10	1	14	1	19	1
0.5	0.7	1	1	5	1	9	1	9	1	17	1
0.6	0.7	2	1	7	1	12	1	17	1	22	1

Table 3 showed the results when different sample sizes when $\alpha=0.10$ the power of the test approximately one and the number of reject between 1 to 22 with different sample sizes and different P_0 and P_1 .

4. Conclusion

This paper proposed two suggested mathematical programming model to determinate power of a statistical test and the critical value K . the two suggested mathematical programming models solved with different sample sizes and different level of significant. For the first suggested model solutions appear as 0-1 integer programming problems. It determined critical or rejection region as points. It rejected some points and accepted some points. That can be difficult to determine the critical or rejection region exactly and the results are difficult to interpret clearly. The results of simulation study for second suggested model when five samples sizes are used with different α , P_0 and P_1 . When the suggested model is applied the power of test is approximant one and the power of test increased when α increased. Also the value of K decreased when α increased. The model solutions appear as integer programming problems but it determined critical or rejection region. The suggested model calculates the critical values which determine the rejection region exactly and the results are easy to interpret clearly. Then the results for the suggested mathematical programming model make balance between the power and the significance.

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